Section 2: Least Squares

Least squares regression:

\[
\min_{\beta} \sum_{n} (y_n - \beta x_n)^2
\]

"objective" or "error" function

In the space of measurements:

\[ \hat{\beta} = \arg \min_{\beta} \sum_{n} (y_n - \beta x_n)^2 \]
can solve this with calculus...

... or, with linear algebra!

\[ \min_{\beta} \sum_{n} (y_n - \beta x_n)^2 \]

\[ \min_{\beta} || \bar{y} - \beta \bar{x} ||^2 \]

**Geometry:**

Note: this is a 2-D cartoon of the N-D vectors, not the two-dimensional \((x,y)\) measurement space of previous plots!

Note: partition of sum of squared data values:

\[ || \bar{y} ||^2 = || \beta_{opt} \bar{x} ||^2 + || \bar{y} - \beta_{opt} \bar{x} ||^2 \]

**Multiple regression:**

\[ \min_{\beta} || \bar{y} - \sum_k \beta_k \bar{x}_k ||^2 = \min_{\beta} || \bar{y} - X \beta ||^2 \]

2D example:
Solution via the “Orthogonality Principle”:

Construct matrix \( X \), containing columns \( \bar{x}_1 \) and \( \bar{x}_2 \)

Orthogonality: \( X^T (\bar{y} - X \bar{\beta}) = 0 \)

Alternatively, can solve using SVD...

\[
\min_{\bar{\beta}} ||\bar{y} - X \bar{\beta}||^2 = \min_{\bar{\beta}} ||\bar{y} - USV^T \bar{\beta}||^2 \\
= \min_{\bar{\beta}} ||U^T \bar{y} - SV^T \bar{\beta}||^2 \\
= \min_{\bar{\beta}^*} ||\bar{y}^* - S \bar{\beta}^*||^2 \\
\text{where } \bar{y}^* = U^T \bar{y}, \quad \bar{\beta}^* = V^T \bar{\beta}
\]

Solution: \( \beta^*_{\text{opt},k} = y_k^*/s_k \), for each \( k \)

or \( \bar{\beta}^*_{\text{opt}} = S^\# \bar{y}^* \implies \bar{\beta}_{\text{opt}} = VS^\# U^T \bar{y} \)

[on board: transformations, elliptical geometry]

Fitting a parametric model (general)

Experimental Data \( \bar{x}_n \)  \( \rightarrow \) Model \( f_\beta(\bar{x}) \) \( \rightarrow \) \( \bar{y}_n \)

To fit model \( f_\beta(\bar{x}) \) to data \( \{\bar{x}_n, \bar{y}_n\} \),

optimize parameters \( \beta \) to minimize an error function:

\[
\min_{\beta} \sum_n E(\bar{y}_n, f_\beta(\bar{x}_n))
\]

Ingredients: data, model, error function, optimization method
Optimization

- Heuristics, exhaustive search, (pain & suffering)
- Iterative descent, (possibly) nonunique
- Iterative descent, guaranteed
- Closed-form guaranteed

Convex
Quadratic
Smooth ($C^2$)

Interpretation warning: fitting a line does not guarantee data actually lie along a line

These 4 data sets give the same regression fit, and same error:

(need a better model)
(need a better error metric)
(need a better experiment)

[Anscombe, 1973]

Polynomial regression

Observation
\( \bar{y} \)  \( \bar{x}_0 \)  \( \bar{x}_1 \)  \( \bar{x}_2 \)

\(-\beta_0\)  \(-\beta_1\)  \(-\beta_2\)
Polynomial regression - how many terms?

(to be continued, when we get to “statistics”...)

Weighted Least Squares

\[
\min_{\beta} \sum_n \left( w_n (y_n - \beta x_n) \right)^2
\]

\[
= \min_{\beta} \left\| W \hat{y} - \beta \hat{x} \right\|^2
\]

diagonal matrix

Solution via simple extensions of basic regression solution
(i.e., let \( \hat{y}^* = W \hat{y} \) and \( \hat{x}^* = W \hat{x} \) and solve for \( \beta \))

Outliers
Outliers

"Trimming"… discard points with large error.
Note: a special case of weighted least squares.

Trimming can be done iteratively (discard outlier, re-fit, repeat),
a so-called "greedy" method. When do you stop?
More generally, use a “robust” error metric. For example:

\[ f(d) = d^2 \]

\[ f(d) = \log(c^2 + d^2) \]

“Lorentzian”

Note: generally can’t obtain solution directly (i.e., requires an iterative optimization procedure). In some cases, can use iteratively re-weighted least squares (IRLS):

**Iteratively Re-weighted Least Squares (IRLS)**

initialize: \( w_n^{(0)} = 1 \)

\[
\beta^{(i)} = \arg \min_{\beta} \sum_n \omega_n^{(i)} (y_n - \beta x_n)^2
\]

\[
\omega_n^{(i+1)} = \left| \frac{f'(y_n - \beta^{(i)} x_n)}{y_n - \beta^{(i)} x_n} \right|
\]

(one of many variants)

**Constrained Least Squares**

Linear constraint:

\[
\arg \min_{\tilde{\beta}} \left\| \tilde{y} - X \tilde{\beta} \right\|^2, \quad \text{where} \quad \tilde{c}^T \tilde{\beta} = 1
\]

Quadratic constraint:

\[
\arg \min_{\beta} \left\| X \beta \right\|^2, \quad \text{where} \quad \left\| \beta \right\|^2 = 1
\]

Can be solved exactly using linear algebra (SVD)...
Standard Least Squares regression

Error is vertical distance (in the “dependent variable”) from the fitted line...

\[
\arg\min_{\beta} ||\tilde{y} - \beta \tilde{x}||^2
\]
Total Least Squares Regression
(a.k.a “orthogonal regression”)

Error is squared distance from the fitted line...

expressed as: \( \min_{\hat{u}} \|D\hat{u}\|^2, \) where \( \|\hat{u}\|^2 = 1 \)

Note: “data” matrix \( D \) now includes both \( x \) and \( y \) coordinates

Variance of data \( D \), projected onto axis \( \hat{u} \):
\[
||USV^T\hat{u}||^2 = ||SV^T\hat{u}||^2 = ||S\hat{u}^*||^2 = ||\hat{u}^{**}||^2,
\]
where \( D = USV^T, \hat{u}^* = V^T\tilde{u}, \) \( \hat{u}^{**} = S\tilde{u}^* \)

Set of \( \hat{u} \)'s of length 1 (i.e., unit vectors)
Set of \( \hat{u}^* \)'s of length 1 (i.e., unit vectors)
First two components of \( \hat{u}^{**} \) (rest are zero!),
for three example \( S \)'s.

Descriptive statistics

Data
“Dispersion”

“Central tendency”
Descriptive statistics: **Central tendency**

- We often summarize data with averages. Why?
- Average minimizes the squared error (as in regression!):
  \[
  \bar{x} = \arg\min_c \frac{1}{N} \sum_{n=1}^{N} (x_n - c)^2 = \frac{1}{N} \sum_{n=1}^{N} x_n = \frac{1}{N} \mathbf{1}^T \mathbf{x}
  \]
  
  - In general: minimize \(L_p\) norm: \(\arg\min_c \left[ \frac{1}{N} \sum_{n=1}^{N} |x_n - c|^p \right]^{1/p}\)
  - \(p = 1\) : median, \(m_x\)
  - \(p \to 0\) : mode (location of maximum)
  - \(p \to \infty\) : midpoint of range
- Issues: outliers, asymmetry, bimodality

Descriptive statistics: **Dispersion**

- Sample variance (squared standard deviation):
  \[
  s^2 = \min_c \frac{1}{N} \sum_{n=1}^{N} (x_n - c)^2 = \frac{1}{N} \sum_{n=1}^{N} (x_n - \bar{x})^2
  \]
  \[
  = \frac{1}{N} \sum_{n=1}^{N} x_n^2 - \bar{x}^2 = \frac{1}{N} \|\mathbf{x}\|^2 - \bar{x}^2
  \]

  Mean absolute deviation (MAD) about the median:
  \[
  d_e = \frac{1}{N} \sum_{n=1}^{N} |x_n - m_x|
  \]
  
  - Quantiles (eg: “90% of data lie in range [1.5 8.2]“)

Descriptive statistics: **Multi-D**

Data points: \(\{ \mathbf{d}_n \}_n \in [1...N]\), in 2-D: \(\mathbf{d}_n = [x_n \ y_n]\)

As in 1D: define central tendency as vector that minimizes the sum of squared distances to all data points:

\[
\mathbf{d} \equiv \arg\min_{\mathbf{c}} \sum_n \| \mathbf{d}_n - \mathbf{c} \|^2
\]

\[
= \arg\min_{c_x, c_y} \sum_n (x_n - c_x)^2 + (y_n - c_y)^2 \quad \text{(in 2-D)}
\]

\[
= \frac{1}{N} \sum_{n=1}^{N} \mathbf{d}_n = \begin{bmatrix} \bar{x} \\ \bar{y} \end{bmatrix}
\]
**Descriptive statistics: Multi-D**

Data points: \( \{ \vec{d}_n \} \) \( n \in [1...N] \)

Sample mean (average):
\[
\bar{d} = \frac{1}{N} \sum_{n=1}^{N} d_n = \left[ \bar{x} \bar{y} \right]
\]

Sample covariance:
\[
C_d = \frac{1}{N} \sum_{n=1}^{N} (\vec{d}_n - \bar{d})(\vec{d}_n - \bar{d})^T = \frac{1}{N} \sum_{n=1}^{N} \vec{d}_n \vec{d}_n^T - \bar{d} \bar{d}^T
\]
\[
= \frac{1}{N} \begin{bmatrix} ||\vec{x}||^2 & \vec{x}^T \vec{y} \\ \vec{y}^T \vec{x} & ||\vec{y}||^2 \end{bmatrix} - \begin{bmatrix} \bar{x}^2 & \bar{x} \bar{y} \\ \bar{y} \bar{x} & \bar{y}^2 \end{bmatrix} = \begin{bmatrix} s_x^2 & s_{xy} \\ s_{xy} & s_y^2 \end{bmatrix}
\]

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**Total Least Squares Regression**
(a.k.a. “orthogonal regression”)

Error is squared distance from the fitted line...

expressed as: \( \min_{\vec{u}} ||D\vec{u}||^2 \), where \( ||\vec{u}||^2 = 1 \)

Note: “data” matrix \( D \) now includes both \( x \) and \( y \) coordinates

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**Olympic gold medalists** (Rio, 2016)

- Michelle Carter (USA)
- Sandra Perković (Croatia)
- Thomas Röhler (Germany)

3D geometry: Javelin, Discus, Shotput…
Principal Component Analysis (PCA)

The shape of a data cloud can be summarized with an ellipse (ellipsoid), centered around the mean, using a simple procedure:

1. Subtract mean of all data points, to re-center around origin
2. Assemble centered data vectors in rows of a matrix, \( D \)
3. Compute the SVD:
   \[
   D = U S V^T
   \]
   or just use the smaller matrix
   \[
   C = D^T D = V S^T S V^T = V A V^T
   \]
4. Columns of \( V \) are the principal components (axes) of the ellipsoid, diagonal elements \( s_k \) or \( \sqrt{\lambda_k} \) are the corresponding principle radii, and their product is the volume.

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Example: PCA for dimensionality reduction and visualization

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Fig 1. Relating biological and model networks using population analyses: Because a model network typically does not attempt to replicate the precise anatomical connectivity of a biological network, there is not a one-to-one correspondence of each biological neuron with a model neuron. Dimensionality reduction can be used to obtain a concise summary of the population activity from each network. This provides common ground for incisive comparisons between biological and model networks. Discrepancies in the population activity structure between biological and model networks can then help to refine model networks.

[Williamson, Doiron, Smith, Yu 2019]

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[Russo et. al., 2018]
Eigenvectors/eigenvalues

- An eigenvector of a matrix is a vector that is rescaled by the matrix (i.e., the direction is unchanged).
- The corresponding scale factor is called the eigenvalue.
- For matrix $C = D^T D = V \Lambda V^T$ the columns of $V$ (denoted $\hat{v}_k$) are eigenvectors, with corresponding eigenvalues $\lambda_k$:

$$C \hat{v}_k = V \Lambda V^T \hat{v}_k = V \Lambda \hat{e}_k = \lambda_k \hat{v}_k$$